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0	0	0	0	0	0	0	0	-8	0	0	0	0	0	0	0
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1	-1	-1	-1	-1	-1	-1	-1	-1	-1	-1	-1	-1	-1	-1	-1

Statistics

Lecture 7

a general introduction to inference.

MATH 414

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We focus on this case mostly.

Examples.

① $X \sim \text{Exp}(\theta)$ but we don't know θ .

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- ③ $X \sim \text{Gamma}(\alpha, \beta)$, both α and β are unknown.
- ④ $X \sim \text{Normal}(\mu, \sigma^2)$, both μ and σ^2 are unknown

We write $f(x; \theta)$ or $p(x; \theta)$ where $\theta \in \Omega$ for a specified set Ω .

The quantity θ is called a parameter of X .

One part of statistics is to estimate parameters.

Estimation is done through samples. Let X_1, X_2, \dots, X_n be i.i.d. r.v.s. Then they constitute a sample of size n from the common distribution X .

Let X_1, X_2, \dots, X_n denote a sample on a random variable X . Let $T = T(X_1, X_2, \dots, X_n)$ be a function of the sample. Then T is called a statistic.

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Once the sample is drawn, then t is called the realization of T , where $t = T(x_1, x_2, \dots, x_n)$ and x_1, x_2, \dots, x_n is the **realization of the sample**.

Point Estimators

The function T of the random sample X_1, X_2, \dots, X_n is called a point estimator for θ . A realization t of T is called an estimate.

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Let x_1, x_2, \dots, x_n be a sample on an r.v. X with p.d.f. $f(x, \theta)$, $\theta \in \Omega$. Let $T = T(x_1, x_2, \dots, x_n)$ be a statistic. We say that T is unbiased estimator of θ if $E[T] = \theta$.

maximum likelihood Estimator (m.l.e)

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The joint distribution of the random sample X_1, \dots, X_n is written as

$$\prod_{i=1}^n f(x_i; \theta)$$

Consider this as a function of θ

↪ since it is i.i.d.

$$L(\theta) = L(\theta; x_1, x_2, \dots, x_n) = \prod_{i=1}^n f(x_i; \theta)$$

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This is called the likelihood function of the random sample.

In order to estimate θ , we might want to find the parameter θ that maximizes

$L(\theta)$

$$\hat{\theta} = \text{Argmax } L(\theta)$$

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sometimes one works with the log of the likelihood function since log function is monotone increasing.

We check for maximality

by checking $\frac{\partial l(\theta)}{\partial \theta} = 0$ for $l(\theta) = \log L(\theta)$

Examples

① Exponential Distribution

Suppose a random sample x_1, x_2, \dots, x_n has the

identical distribution $f(x) = \theta^{-1} e^{-x/\theta}$

$$L(\theta) = \prod_{i=1}^n f(x_i; \theta) = \theta^{-n} e^{-\sum_{i=1}^n x_i/\theta}$$

$$\therefore \ell(\theta) = \log L(\theta) = -n \log \theta - \theta^{-1} \sum_{i=1}^n x_i$$

$$\frac{\partial \ell}{\partial \theta}(\theta) = -\frac{n}{\theta} + \frac{1}{\theta^2} \sum_{i=1}^n x_i = 0 \Rightarrow \theta = \frac{\sum_{i=1}^n x_i}{n}$$

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m.l.e is sample mean which is unbiased.

Examples

① Normal Distribution Let X have a $N(\mu, \sigma^2)$ distribution. In this case $\underline{\theta} = (\mu, \sigma)$ is a vector.

The m.l. function

$$L(\underline{\theta}) = L(\mu, \sigma) = \prod_{i=1}^n \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{(x_i - \mu)^2}{2\sigma^2}\right)$$

$$\therefore \ell(\mu, \sigma) = -\frac{n}{2} \log 2\pi - n \log \sigma - \frac{1}{2} \sum_{i=1}^n \left(\frac{x_i - \mu}{\sigma}\right)^2.$$

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$$\Rightarrow \hat{\mu} = \bar{x} \quad \sim \text{unbiased}$$

$$\hat{\sigma} = \frac{\sum (x_i - \bar{x})^2}{n} \quad \sim \text{biased}$$

Exercises for assignment

- ① m.l.e. for binomial,
uniform distributions etc.